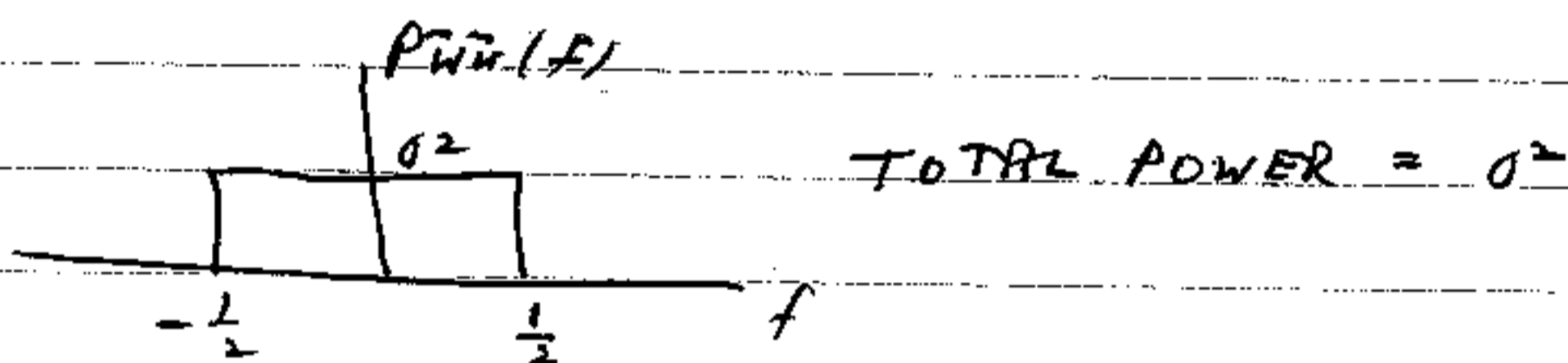


$$\begin{aligned}
 &= E(W_R[n] W_R[n+k]) + E(W_I[n] W_I[n+k]) \\
 &= \sigma^2/2 \delta[k] + \sigma^2/2 \delta[k] \\
 &= \sigma^2 \delta[k]
 \end{aligned}$$

$$\begin{aligned}
 P_{\tilde{w}\tilde{w}}(f) &= \mathcal{F}\{r_{\tilde{w}\tilde{w}}(k)\} \\
 &= \sigma^2 \quad |f| \leq \frac{1}{2}
 \end{aligned}$$



COMPLEX GAUSSIAN PDF

WE KNOW PDF OF $\{W_R[0], \dots, W_R[N-1],$
 $W_I[0], \dots, W_I[N-1]\}$

SINCE THEY ARE IID WITH PDF $N(0, \sigma^2/2)$

EXAMPLE: N=1

$$\begin{aligned}
 p(W_R[0], W_I[0]) &= p(W_R[0]) p(W_I[0]) \\
 &= \frac{1}{\sqrt{2\pi \sigma^2/2}} e^{-\frac{1}{2\sigma^2/2} W_R^2[0]} \\
 &\quad \cdot \frac{1}{\sqrt{2\pi \sigma^2/2}} e^{-\frac{1}{2\sigma^2/2} W_I^2[0]}
 \end{aligned}$$

$$= \frac{1}{\pi \sigma^2} e^{-\frac{1}{\sigma^2} (W_R^2[0] + W_I^2[0])}$$

$$= \frac{1}{\pi \sigma^2} e^{-\frac{1}{\sigma^2} |\tilde{W}[0]|^2}$$

COMPLEX
↓
NUMBERS

DEFINED FOR ALL $\tilde{W}[0] \in \mathbb{C}$

THIS PDF $p(W_R[0], W_I[0]) = p(\tilde{W}[0])$

CALLED THE COMPLEX GAUSSIAN PDF
OF THE COMPLEX RANDOM VARIABLE $\tilde{W}[0]$

ALSO WE SAY $\tilde{W}[0] \sim CN(0, \sigma^2)$

NOTE

$$\text{VAR}(\tilde{W}[0]) = E(|\tilde{W}[0]|^2) - |E(\tilde{W}[0])|^2$$

$$E(\tilde{W}[0]) = E(W_R[0] + jW_I[0])$$

$$= E(W_R[0]) + jE(W_I[0]) = 0$$

$$\text{VAR}(\tilde{W}[0]) = E(|\tilde{W}[0]|^2) = \sigma^2$$

WHY?

FOR CWGN SHOW THAT

$$p(\tilde{W}[0], \dots, \tilde{W}[N-1]) = \frac{1}{(\pi \sigma^2)^N} e^{-\frac{1}{\sigma^2} \sum_{n=0}^{N-1} |\tilde{W}[n]|^2}$$

OR

$$p(\underline{\tilde{w}}) = \frac{1}{(\pi\sigma^2)^N} e^{-\frac{1}{\sigma^2} \underline{\tilde{w}}^H \underline{\tilde{w}}} \quad (A = \mathbf{1}^T)$$

NOTE SIMILARITY TO REAL WGN PDF

$$p(\underline{w}) = \frac{1}{(2\pi\sigma^2)^N} e^{-\frac{1}{2\sigma^2} \underline{w}^T \underline{w}}$$

COMPLEX DATA DETECTORS

(CHAPTER 13)

WISH TO EXTEND MATCHED FILTER
AND ESTIMATOR-CORRELATORMATCHED FILTERKNOWN COMPLEX DETERMINISTIC SIGNAL
IN CWGN

$$\left. \begin{array}{l} H_0: \tilde{x}(n) = \tilde{w}(n) \\ H_1: \tilde{x}(n) = \tilde{s}(n) + \tilde{w}(n) \end{array} \right\} n=0, 1, \dots, N-1$$

\uparrow \uparrow CWGN
 KNOWN $\sim CN(0, \sigma^2)$

USING NP AND GAUSSIAN PDF FOR CWGN

$$L(\tilde{\underline{x}}) = \frac{p(\tilde{\underline{x}}; \mathcal{H}_1)}{p(\tilde{\underline{x}}; \mathcal{H}_0)}$$

(RECALL THAT $p(\tilde{\underline{x}})$ IS REALLY
 $p(\tilde{x}_1(0), \dots, \tilde{x}_1(N-1)) = p(x_R(0), x_I(0), \dots,$
 $x_R(N-1), x_I(N-1))$

USING COMPLEX NOTATION SIMPLIFIES
 ALGEBRA - NO NEW THEORY REQUIRED,
 CAN ALSO VIEW APPROACH AS
 MAPPING FROM $\begin{pmatrix} x \\ y \end{pmatrix}$ TO $x + jy$ OR
 $\mathbb{R}^2 \rightarrow \mathbb{C}$)

UNDER \mathcal{H}_0 $\tilde{\underline{x}} = \tilde{\underline{w}} \sim \mathcal{CN}(\underline{0}, \sigma^2 \mathbf{I})$

$$p(\tilde{\underline{w}}) = \frac{1}{\pi^N \sigma^{2N}} e^{-\frac{1}{\sigma^2} \tilde{\underline{w}}^H \tilde{\underline{w}}}$$

UNDER \mathcal{H}_1 $\tilde{\underline{x}} = \tilde{\underline{s}} + \tilde{\underline{w}} \sim \mathcal{CN}(\tilde{\underline{s}}, \sigma^2 \mathbf{I})$

$$p(\tilde{\underline{x}}) = \frac{1}{\pi^N \sigma^{2N}} e^{-\frac{1}{\sigma^2} (\tilde{\underline{x}} - \tilde{\underline{s}})^H (\tilde{\underline{x}} - \tilde{\underline{s}})}$$

$$\begin{aligned} \Rightarrow \ln L(\tilde{\underline{x}}) &= -\frac{1}{\sigma^2} \left((\tilde{\underline{x}} - \tilde{\underline{s}})^H (\tilde{\underline{x}} - \tilde{\underline{s}}) - \tilde{\underline{x}}^H \tilde{\underline{x}} \right) \\ &= -\frac{1}{\sigma^2} \left[-\tilde{\underline{x}}^H \tilde{\underline{s}} - \tilde{\underline{s}}^H \tilde{\underline{x}} + \tilde{\underline{s}}^H \tilde{\underline{s}} \right] \end{aligned}$$

$$= \frac{2}{\sigma^2} \operatorname{Re}(\tilde{y}^H \tilde{x}) - \frac{1}{\sigma^2} \tilde{y}^H \tilde{y}$$

OR WE DECIDE \mathcal{H}_1 IF

$$T(\tilde{x}) = \operatorname{Re}(\tilde{y}^H \tilde{x})$$

$$= \operatorname{Re}\left(\sum_{n=0}^{N-1} \tilde{x}[n] \tilde{y}^*[n]\right) > \gamma'$$

EXAMPLE: $\tilde{y}[n] = \tilde{A}$ COMPLEX
DC LEVEL

$$T(\tilde{x}) = \operatorname{Re}\left(\sum_{n=0}^{N-1} \tilde{x}[n] \tilde{A}^*\right)$$

$$\text{UNDER } \mathcal{H}_0 \quad T(\tilde{x}) = \operatorname{Re}\left(\sum_n \tilde{w}[n] \tilde{A}^*\right)$$

AND $E(T) = 0$ SHOW THIS.

$$\text{UNDER } \mathcal{H}_1 \quad T(\tilde{x}) = \operatorname{Re}\left(\sum_n (\tilde{A} + \tilde{w}[n]) \tilde{A}^*\right)$$

$$= \operatorname{Re}\left(\sum_n |\tilde{A}|^2 + \sum_n \tilde{w}[n] \tilde{A}^*\right)$$

$$= \underbrace{\sum_n |\tilde{A}|^2}_{\text{ENERGY} = \mathcal{E}} + \operatorname{Re}\left(\sum_n \tilde{w}[n] \tilde{A}^*\right)$$

CAN SHOW THAT

$$P_{FA} = Q\left(\frac{\gamma'}{\sqrt{\sigma^2 \mathcal{E}/2}}\right)$$

$$\mathcal{E} = \sum_{n=0}^{N-1} |\tilde{y}[n]|^2$$

$$P_D = Q\left(\frac{\delta' - \epsilon}{\sqrt{\sigma^2 \epsilon / 2}}\right)$$

$$\text{OR } P_D = Q\left(Q^{-1}(1 - \text{FA}) - \sqrt{d^2}\right)$$

$$\text{WHERE } d^2 = \frac{2\epsilon}{\sigma^2}$$

d^2 IS TWICE WHAT WE HAD FOR REAL CASE - DUE TO MEAN UNDER \mathcal{H}_1 IS REAL \rightarrow DISCARD IMAGINARY PART OF $\sum_n x[n] \tilde{x}^*[n]$ TO FIND $T(\tilde{x})$

ESTIMATOR-CORRELATOR

RECALL FOR CWGN

$$p(\tilde{x}) = \frac{1}{(2\pi)^N \sigma^{2N}} e^{-\frac{1}{\sigma^2} \tilde{x}^H \tilde{x}}$$

$$\text{AND } \underline{C}_{\tilde{x}} = E(\tilde{x} \tilde{x}^H) = \sigma^2 \underline{I}$$

TO SEE THIS CONSIDER $N=2$:

$$\begin{aligned} \underline{C}_{\tilde{x}} &= E\left(\begin{bmatrix} \tilde{x}[0] \\ \tilde{x}[1] \end{bmatrix} \begin{bmatrix} \tilde{x}^*[0] & \tilde{x}^*[1] \end{bmatrix}\right) \\ &= \begin{bmatrix} E(|\tilde{x}[0]|^2) & E(\tilde{x}[0] \tilde{x}^*[1]) \\ E(\tilde{x}[1] \tilde{x}^*[0]) & E(|\tilde{x}[1]|^2) \end{bmatrix} \end{aligned}$$

ALREADY SHOWED $E(|\tilde{x}[n]|^2) = \text{VAR}(\tilde{x}[n]) = \sigma^2$

$$E(\tilde{x}(l_0) \tilde{x}^*(l_1)) = E((x_R(l_0) + jx_I(l_0)) \\ (x_R(l_1) - jx_I(l_1)))$$

$$= E(x_R(l_0)x_R(l_1)) - jE(x_R(l_0)x_I(l_1)) \\ + jE(x_I(l_0)x_R(l_1)) + E(x_I(l_0)x_I(l_1))$$

= 0 SINCE ALL RANDOM VARIABLES
ARE INDEPENDENT

$$\Rightarrow p(\tilde{\mathbf{x}}) = \frac{1}{\pi^N \text{DET}(\mathbf{C}_{\tilde{\mathbf{x}}})} e^{-\tilde{\mathbf{x}}^H \mathbf{C}_{\tilde{\mathbf{x}}}^{-1} \tilde{\mathbf{x}}}$$

THIS IS CALLED THE COMPLEX
MULTIVARIATE GAUSSIAN PDF AND
WE USE $\tilde{\mathbf{x}} \sim \mathcal{CN}(\mathbf{0}, \mathbf{C}_{\tilde{\mathbf{x}}})$

SEE KAY-II, 1993 FOR DETAILS

BACK TO DETECTION PROBLEM -

CONSIDER A COMPLEX RANDOM SIGNAL
WITH COVARIANCE MATRIX $\mathbf{C}_{\tilde{\mathbf{x}}}$ (AND
ZERO MEAN) AND INDEPENDENT OF
CWGN

$$\left. \begin{array}{l} \mathcal{H}_0: \tilde{x}(l) = \tilde{w}(l) \\ \mathcal{H}_1: \tilde{x}(l) = \tilde{s}(l) + \tilde{w}(l) \end{array} \right\} l = 0, 1, \dots, N-1$$

\uparrow \uparrow
 RANDOM CWGN
 SIGNAL

OR $\tilde{\underline{x}} \sim \text{CN}(\underline{0}, \sigma^2 \underline{I})$ UNDER \mathcal{H}_0
 $\text{CN}(\underline{0}, \underline{C}\hat{\underline{s}} + \sigma^2 \underline{I})$ UNDER \mathcal{H}_1

NP DETECTOR DECIDES \mathcal{H}_1 IF

$$L(\tilde{\underline{x}}) = \frac{p(\tilde{\underline{x}}; \mathcal{H}_1)}{p(\tilde{\underline{x}}; \mathcal{H}_0)} > \delta$$

WHERE

$$p(\tilde{\underline{x}}; \mathcal{H}_0) = \frac{1}{\pi^N \sigma^{2N}} e^{-\frac{1}{\sigma^2} \tilde{\underline{x}}^H \tilde{\underline{x}}}$$

$$p(\tilde{\underline{x}}; \mathcal{H}_1) = \frac{1}{\pi^N \text{DET}(\underline{C}\hat{\underline{s}} + \sigma^2 \underline{I})} e^{-\tilde{\underline{x}}^H (\underline{C}\hat{\underline{s}} + \sigma^2 \underline{I})^{-1} \tilde{\underline{x}}}$$

RESULT: DECIDE \mathcal{H}_1 IF

$$T(\tilde{\underline{x}}) = \tilde{\underline{x}}^H \underbrace{\underline{C}\hat{\underline{s}} (\underline{C}\hat{\underline{s}} + \sigma^2 \underline{I})^{-1}}_{\hat{\underline{s}} = \text{MMSE ESTIMATOR OF } \underline{s}} \tilde{\underline{x}} > \delta'$$

EXAMPLE: NON FLUCTUATING POINT TARGET

$$\tilde{\underline{s}}(n) = \tilde{\underline{A}} \tilde{\underline{h}}(n)$$

↑ ↑
RANDOM KNOWN

ASSUME $\tilde{A} \sim \text{CN}(0, \sigma_A^2)$

IF $\tilde{h}[n] = e^{j2\pi f_D n}$
 \uparrow DOPPLER FREQ.

$\tilde{J}[n] = |\tilde{A}| e^{j(2\pi f_D n + \angle \tilde{A})}$
 \uparrow RAYLEIGH RANDOM VAR. \uparrow UNIFORM RANDOM VAR.

INDEPENDENT

SINCE $|\tilde{A}| = \sqrt{A_R^2 + A_I^2}$
 $\angle \tilde{A} = \arctan \frac{A_I}{A_R}$

AND $A_R \sim N(0, \sigma_A^2/2)$ > INDEPENDENT
 $A_I \sim N(0, \sigma_A^2/2)$

THIS IS STANDARD MODEL FOR COMPLEX
 ACTIVE SONAR / RADAR RETURN FROM
 NONFLUCTUATING POINT TARGET.

\uparrow NO BANDWIDTH \uparrow NO DELAY SPREAD
 EXPANSION (EXCEPT DOPPLER SHIFT)

$\tilde{h}(t, \tau) = \tilde{A} \delta(\tau)$

\uparrow DOPPLER SPREAD \uparrow DELAY SPREAD

LINEAR TIME-VARYING
 RANDOM CHANNEL/
 TARGET IMPULSE
 RESPONSE)

WHAT IS $\underline{C}_{\tilde{S}}$?

$$\underline{\tilde{S}} = \tilde{A} \tilde{h}$$

$$\underline{C}_{\tilde{S}} = E(\underline{\tilde{S}} \underline{\tilde{S}}^H) = E(\tilde{A} \tilde{h} \tilde{A}^* \tilde{h}^H)$$

$$= E(|\tilde{A}|^2) \underline{h} \underline{h}^H = \sigma_A^2 \underline{h} \underline{h}^H$$

THIS IS RANK ONE \Rightarrow CAN USE
WOODBURY TO INVERT $(\underline{C}_{\tilde{S}} + \sigma^2 \underline{I})$

RESULT : DECIDE \mathcal{H}_1 IF

$$T'(\underline{\tilde{x}}) = \left| \sum_{n=0}^{N-1} \tilde{x}(n) \tilde{h}^*(n) \right|^2 > \gamma''$$

EXAMPLE : $\tilde{h}(n) = e^{j2\pi f_0 n}$

$$\Rightarrow \underline{T'(\underline{\tilde{x}})} = \frac{1}{N} \left| \sum_{n=0}^{N-1} \tilde{x}(n) e^{-j2\pi f_0 n} \right|^2$$

PERIODOGRAM

(f_0 IS NOT REALLY
KNOWN)

SEE BOOK FOR PERFORMANCE