

TO IMPLEMENT GLRT

$$L(x) = \frac{p(x; \hat{A}, H_1)}{p(x; H_0)}$$

NEED MLE  $\hat{A}$ . BUT

$$p(x; A, H_1) = C_1^N e^{-\frac{1}{2} C_2 \sum_{n=0}^{N-1} (x[n] - A)^2}$$

MUST MINIMIZE  $J(A) = \sum_{n=0}^{N-1} (x[n] - A)^2$   
OVER  $A \Rightarrow$  HARD TO DO ANALYTICALLY.

BUT

$$TR(x) = \frac{\left. \frac{\partial \ln p(x; A)}{\partial A} \right|_{A=0}}{I^{-1}(0)}$$

$$\begin{aligned} \left. \frac{\partial \ln p(x; A)}{\partial A} \right|_{A=0} &= -2 C_2 (-1) \sum_n (x[n] - A)^3 \Big|_{A=0} \\ &= 2 C_2 \sum_n x^3[n] \end{aligned}$$

$$TR(x) = \frac{4 C_2^2}{I(0)} \left( \sum_{n=0}^{N-1} x^3[n] \right)^2$$

OR

$$TR(\underline{x}) = \frac{4N^2 C_2^2}{I(0)} \left( \frac{1}{N} \sum_{n=0}^{N-1} x^3(n) \right)^2$$

SEE EX. 6.9 FOR EXACT EXPRESSIONS

NOTE THAT UNDER  $\mathcal{H}_0$ 

$$E(x^3(n)) = 0 \quad \text{WHY?}$$

$$\Rightarrow \frac{1}{N} \sum_{n=0}^{N-1} x^3(n) \approx 0$$

UNDER  $\mathcal{H}_1$  THERE WILL BE A  
SIGNAL CONTRIBUTION OF  $\left( \frac{1}{N} \sum A^2 \right)^2$   
 $= A^6$

CASE 2: NUISANCE PARAMETERS

$$\mathcal{H}_0: \underline{\theta}_r = \underline{\theta}_{r0}, \underline{\theta}_s$$

$$\mathcal{H}_1: \underline{\theta}_r \neq \underline{\theta}_{r0}, \underline{\theta}_s$$

RAO TEST BECOMES

$$TR(\underline{x}) = \frac{\partial \text{LN}P(\underline{x}; \underline{\theta})}{\partial \underline{\theta}_r} \Big|_{\underline{\theta} = \tilde{\underline{\theta}}}^T \left( \underline{I}^{-1}(\tilde{\underline{\theta}}) \right)_{\theta_r \theta_r}$$

$$\cdot \frac{\partial \text{LN}P(\underline{x}; \underline{\theta})}{\partial \underline{\theta}_r} \Big|_{\underline{\theta} = \tilde{\underline{\theta}}}$$

WHERE  $\tilde{\theta} = [\theta_{r0}^T \hat{\theta}_{s0}^T]^T$   
 $=$  MLE OF  $\theta$  UNDER  $H_0$   
 $=$  CONSTRAINED MLE

NEED ONLY FIND MLE UNDER  $H_0 \Rightarrow$

NO SIGNAL PARAMETER

MLE NEEDED

EXAMPLE : DC LEVEL IN WGN

$$H_0: A = 0, \sigma^2 > 0$$

$$H_1: A \neq 0, \sigma^2 > 0$$

$A, \sigma^2$  UNKNOWN, ONLY NEED

MLE UNDER  $H_0$  OR NEED  $\hat{\sigma}^2$

FOR WGN

$$\theta = \begin{pmatrix} A \\ \sigma^2 \end{pmatrix} = \begin{pmatrix} \theta_r \\ \theta_s \end{pmatrix} \quad r=1, s=1$$

$$\theta_{r0} = A_0 = 0$$

$$\hat{\theta}_{s0} = \hat{\sigma}_0^2 = \frac{1}{N} \sum_{n=0}^{N-1} x^2(n) \Rightarrow \tilde{\theta} = \begin{pmatrix} 0 \\ \hat{\sigma}_0^2 \end{pmatrix}$$

RAO TEST BECOMES

$$TR(x) = \frac{\partial \ln p(x; A, \sigma^2)}{\partial A} \bigg|_{\substack{A=0 \\ \sigma^2 = \hat{\sigma}_0^2}} \left[ I^{-1}(A, \sigma^2) \right]_{AA}$$

$$\frac{\partial \ln p(x; A, \sigma^2)}{\partial A} \bigg|_{\substack{A=0 \\ \sigma^2 = \hat{\sigma}_0^2}}$$

WHERE

$$[\underline{I}^{-1}(\tilde{\theta})]_{AA} = \text{AA PARTITION OF } \underline{I}^{-1}(\tilde{\theta})$$

$$\underline{I}(\theta) = \underline{I}(A, \sigma^2)$$

$$= \begin{bmatrix} I_{AA} & I_{A\sigma^2} \\ I_{\sigma^2 A} & I_{\sigma^2\sigma^2} \end{bmatrix}$$

$$= \begin{bmatrix} N/\sigma^2 & 0 \\ 0 & N/2\sigma^4 \end{bmatrix}$$

$$\underline{I}^{-1}(\theta) = \begin{bmatrix} \sigma^2/N & 0 \\ 0 & 2\sigma^4/N \end{bmatrix}$$

$$\underline{I}^{-1}(\tilde{\theta}) = \underline{I}^{-1}(A=0, \sigma^2=\hat{\sigma}_0^2)$$

$$= \begin{bmatrix} \hat{\sigma}_0^2/N & 0 \\ 0 & 2\hat{\sigma}_0^4/N \end{bmatrix}$$

$$[\underline{I}^{-1}(\tilde{\theta})]_{AA} = \hat{\sigma}_0^2/N$$

$$\text{AND } \frac{\partial \ln p(x; A, \sigma^2)}{\partial A} = -\frac{1}{2\sigma^2} \frac{\partial}{\partial A} \sum_n (x/n - A)^2$$

$$= \frac{1}{\sigma^2} \sum_n (x(n) - A)$$

$$\frac{\partial \ln p(x; A, \sigma^2)}{\partial A} \Big|_{A=0, \sigma^2 = \hat{\sigma}_0^2} = \frac{1}{\hat{\sigma}_0^2} \sum x(n)$$

$$= \frac{N \bar{x}}{\hat{\sigma}_0^2}$$

$$TR(\underline{x}) = \left( \frac{N \bar{x}}{\hat{\sigma}_0^2} \right)^2 \frac{\hat{\sigma}_0^2}{N}$$

$$= \frac{N \bar{x}^2}{\hat{\sigma}_0^2}$$

RECALL THAT GLRT WAS

$$2 \ln L_G(\underline{x}) = N \ln \frac{1}{1 - \frac{\bar{x}^2}{\hat{\sigma}_0^2}}$$

TO SHOW THAT THEY ARE

EQUIVALENT AS  $N \rightarrow \infty$  NOTE

THAT AS  $N \rightarrow \infty$  MUST HAVE

$A \rightarrow 0$  (WHY?)

$\Rightarrow \bar{x} \rightarrow 0$  WAY?

$$\bar{x} / \hat{\sigma}_0^2 \ll 1$$

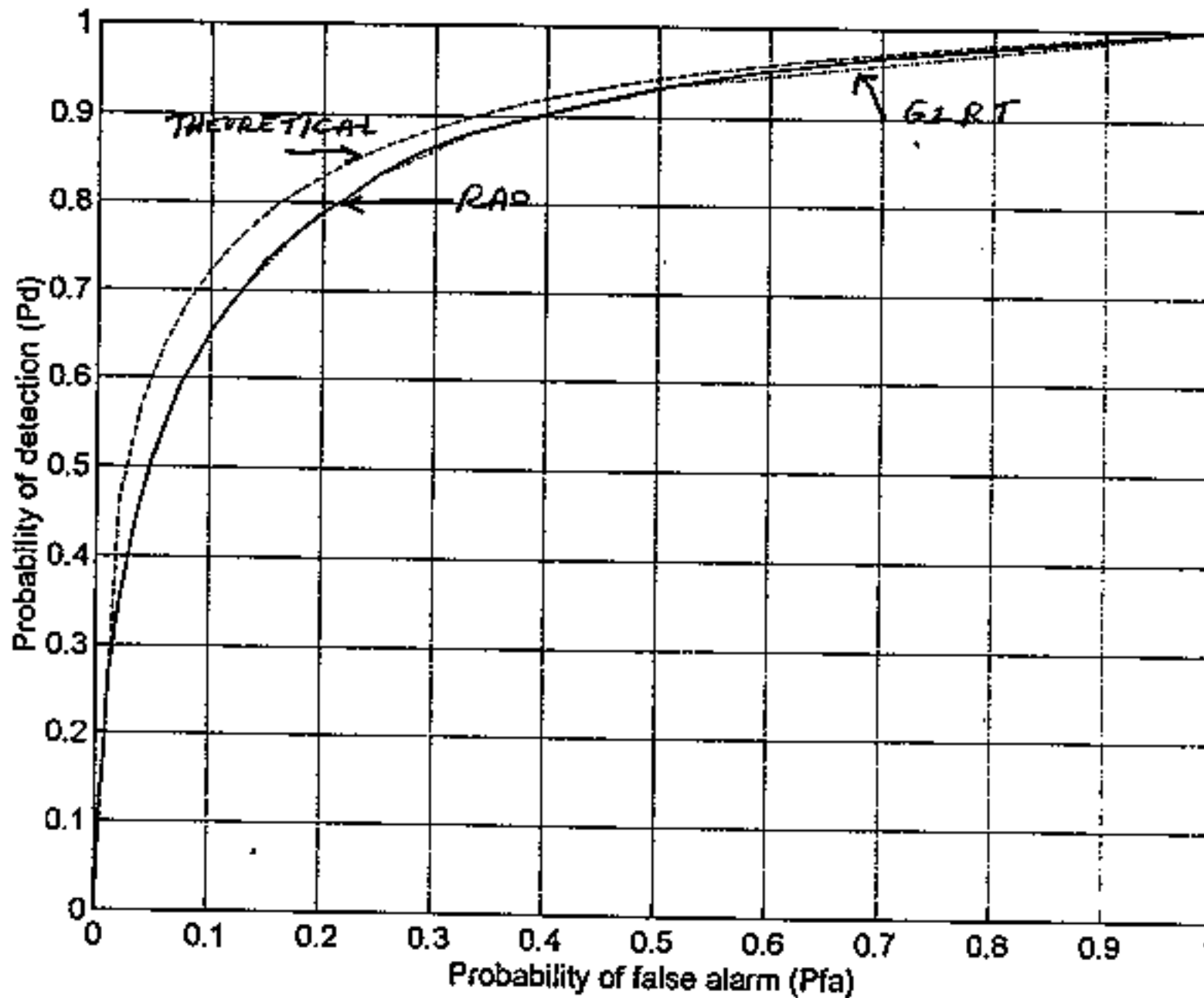
BUT  $\ln \frac{1}{1-x} \approx x$  FOR  $x \ll 1$

$$\Rightarrow 2 \ln L_G(x) \approx N \ln \frac{1}{1 - \frac{\bar{x}^2}{\hat{\sigma}_0^2}}$$

$$\approx N \frac{\bar{x}^2}{\hat{\sigma}_0^2} = \text{TR}(\bar{x})$$

FOR FINITE DATA RECORDS

$$N = 10$$



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% fig67new.m
lambda=5;sig2=1;N=30;A=sqrt(lambda*sig2/N);
nreal=5000;
for i=1:nreal
x0=randn(N,1);x1=x0+A;
y0=(mean(x0))^2/(x0'*x0/N);
y1=(mean(x1))^2/(x1'*x1/N);
TOR(i,1)=N*y0;T1R(i,1)=N*y1;
T0(i,1)=N*log(1/(1-y0));
T1(i,1)=N*log(1/(1-y1));
end
for i=1:51
Pfaa(i,1)=(i-1)/50;
u=Qinv(Pfaa(i)/2);
Pda(i,1)=Q(u-sqrt(lambda))+Q(u+sqrt(lambda));
end
[Pfag,Pdg]=rocurve(T0,T1,51);
[PfaR,PdR]=rocurve(TOR,T1R,51);
plot(Pfag,Pdg,'-.',PfaR,PdR,'-',Pfaa,Pda,'--')
xlabel('Probability of false alarm (Pfa)')
ylabel('Probability of detection (Pd)')
grid
print

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## CHAPTER 7 - DETERMINISTIC SIGNALS - UNKNOWN PARAMETERS

LACK OF SIGNAL KNOWLEDGE  
DEGRADES DETECTOR PERFORMANCE.  
COMPARE SIGNAL KNOWN VS  
SIGNAL UNKNOWN CASES.

- 1) SIGNAL KNOWN  $\Rightarrow$  MATCHED FILTER
- 2) SIGNAL UNKNOWN  $\Rightarrow$  USE GLRT  
 $\Rightarrow$  ENERGY DETECTOR

$$H_0: x(n) = w(n) \quad n = 0, 1, \dots, N-1$$

$$H_1: x(n) = s(n) + w(n) \quad n = 0, 1, \dots, N-1$$

$\uparrow$   
 WGN WITH KNOWN  
 VARIANCE  $\sigma^2$

- 1)  $s(n)$  DETERMINISTIC AND KNOWN  
 $\Rightarrow$  USE  $T(x) = \sum_{n=0}^{N-1} x(n)s(n)$

$$P_D = Q(Q^{-1}(P_{FA}) - \sqrt{E/\sigma^2})$$

- 2)  $s(n)$  DETERMINISTIC AND UNKNOWN



USE A GLRT.

$$L_G(x) = \frac{p(x; \hat{s}[0], \hat{s}[1], \dots, \hat{s}[N-1], H_1)}{p(x; H_0)}$$

TO FIND MLE OF  $s[n]$  MAXIMIZE

$$p(x; s[0], s[1], \dots, s[N-1], H_1) =$$

$$\frac{1}{(2\pi\sigma^2)^{N/2}} e^{-\frac{1}{2\sigma^2} \sum_n (x[n] - s[n])^2}$$

OVER  $s[0], s[1], \dots, s[N-1]$

$$\Rightarrow \hat{s}[n] = x[n]$$

$$L_G(x) = \frac{1}{(2\pi\sigma^2)^{N/2}} \frac{1}{(2\pi\sigma^2)^{N/2}} e^{-\frac{1}{2\sigma^2} \sum x^2[n]}$$

$$\ln L_G(x) = -\frac{1}{2\sigma^2} \sum_n x^2[n]$$

GLRT DECIDES  $H_1$  IF

$$T(x) = \sum_{n=0}^{N-1} x^2[n] > \gamma'$$

QUADRATIC OR  
ENERGY DETECTOR  
(NONCOHERENT)

CAN'T FIND PERFORMANCE FROM PREVIOUS WORK SINCE WE HAD  $J(n) = \text{RANDOM PROCESS}$ .

TO FIND PERFORMANCE ASSUME  $N$  IS LARGE AND APPLY CENTRAL LIMIT THEOREM ( $X(n)$  ARE INDEPENDENT BUT NOT IDENTICALLY DISTRIBUTED)

ASIDE: NONCENTRAL  $\chi^2$  PDF

IF  $X_1, X_2, \dots, X_N$  ARE INDEPENDENT WITH  $X_i \sim N(\mu_i, 1)$ , THEN

$$Y = \sum_{i=1}^N X_i^2 \sim \chi_N^2(\lambda) \quad \begin{array}{l} \leftarrow \text{NONCENTRALITY} \\ \text{PARAMETER} \\ \uparrow \text{DEGREES OF FREEDOM} \end{array}$$

$$\lambda = \sum_{i=1}^N \mu_i^2 \quad \text{SEE CHAPTER 2}$$

$$E(Y) = \lambda + N$$

$$\text{VAR}(Y) = 4\lambda + 2N$$

IF  $\mu_i = 0$  ALL  $i \Rightarrow \lambda = 0$

$\Rightarrow$  GET  $\chi_N^2 = \text{CENTRAL } \chi^2 \text{ PDF}$

TO GET FIRST TWO MOMENTS:

$$T'(\underline{x}) = \frac{T(\underline{x})}{\sigma^2} \sim \begin{matrix} \chi_N^2 & H_0 \\ \chi_N'^2(\lambda) & H_1 \end{matrix}$$

$$\text{WHERE } \lambda = \frac{\sum s^2(n)}{\sigma^2} = \epsilon/\sigma^2$$

$$\Rightarrow E(T; H_0) = E(\sigma^2 \chi_N^2) = N\sigma^2$$

$$\begin{aligned} E(T; H_1) &= E(\sigma^2 \chi_N'^2(\lambda)) \\ &= \sigma^2(\lambda + N) \end{aligned}$$

$$\text{VAR}(T; H_0) = \sigma^4 2N$$

$$\text{VAR}(T; H_1) = \sigma^4(4\lambda + 2N)$$

$$P_{FA} = Q\left(\frac{\delta' - N\sigma^2}{\sqrt{2N\sigma^4}}\right)$$

$$P_D = Q\left(\frac{\delta' - (\lambda + N)\sigma^2}{\sqrt{\sigma^4(4\lambda + 2N)}}\right)$$

$$\delta' = N\sigma^2 + \sqrt{2N\sigma^4} Q^{-1}(P_{FA})$$

$$P_D = Q\left(\frac{\sqrt{2N} Q^{-1}(P_{FA}) - \lambda}{\sqrt{4\lambda + 2N}}\right)$$

AND FOR  $\lambda/N \ll 1$  CAN BE  
SHOWN TO BE (SEE APPENDIX 7A)

$$P_D = Q\left(Q^{-1}(P_{FA}) - \sqrt{\frac{\lambda^2}{2N}}\right)$$

↑  $d_{ED}^2$

$$d_{ED}^2 = \frac{(\mathcal{E}/\sigma^2)^2}{2N}$$

FOR  $\frac{\mathcal{E}}{N} = \frac{\mathcal{E}}{N\sigma^2} \ll 1$

OR  $\frac{\mathcal{E}}{\sigma^2} \ll N$

LOW SNR CASE

FOR MATCHED FILTER

$$d_{MF}^2 = \mathcal{E}/\sigma^2$$

$$\Rightarrow \text{LOSS IN dB} = 10 \log_{10} \frac{d_{MF}^2}{d_{ED}^2} \text{ dB}$$

$$= 10 \log_{10} \frac{2N}{\mathcal{E}/\sigma^2} \text{ dB}$$

THIS IS LOSS DUE TO LACK OF SIGNAL KNOWLEDGE.

EXAMPLE : DC LEVEL IN WGN

(DC LEVEL SIGNAL IS UNKNOWN)

$$\mathcal{E} = NA^2 \Rightarrow \text{FOR } \frac{\mathcal{E}}{\sigma^2} \ll N$$

$$\frac{NA^2}{\sigma^2} \ll N \text{ OR } \frac{A^2}{\sigma^2} \ll 1$$

LOW INPUT SNR CASE

COMPARE REQUIRED INPUT SNR  
FOR GIVEN DETECTION PERFORMANCE

LET  $\eta = A^2/\sigma^2 = \text{INPUT SNR}$

$$d_{MF}^2 = \frac{NA^2}{\sigma^2} = N \eta_{MF}$$

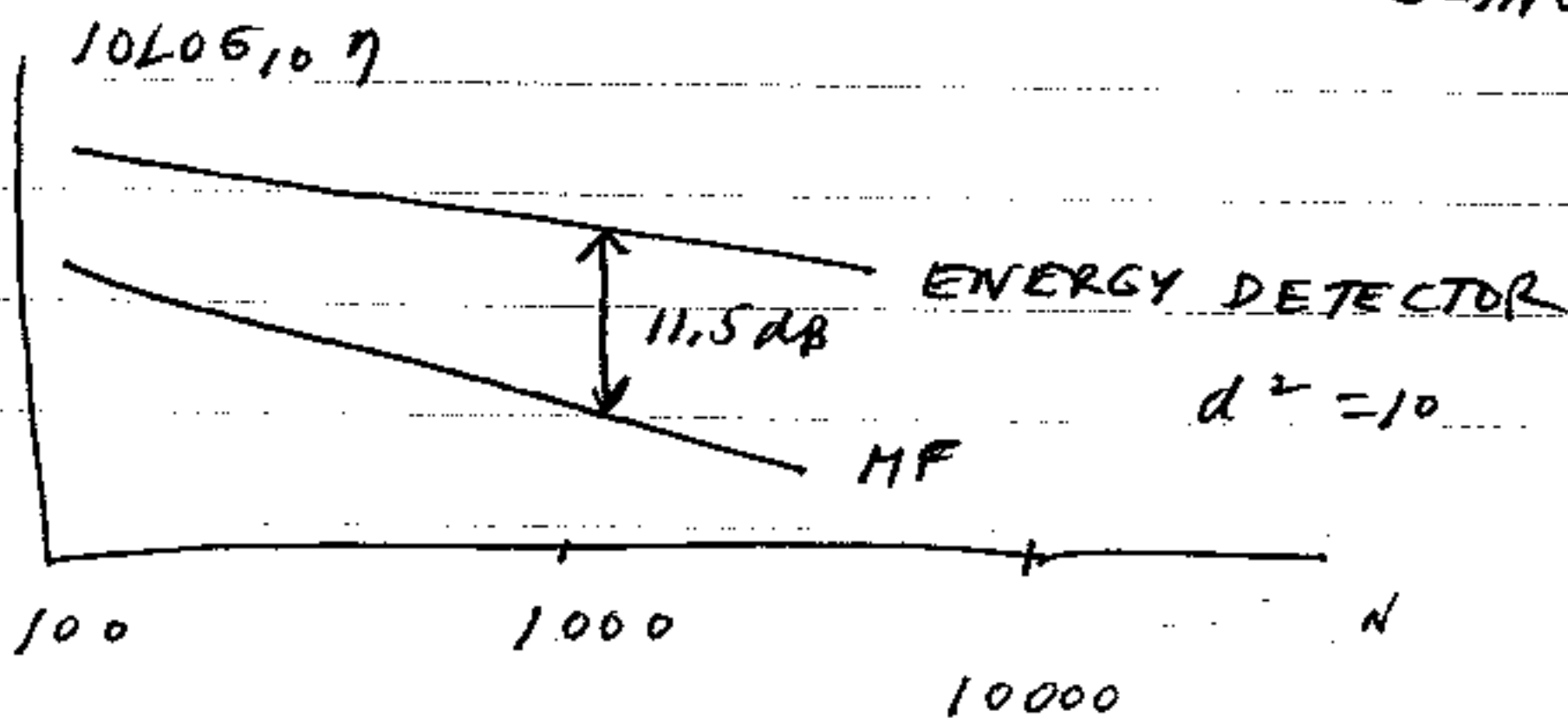
$$d_{ED}^2 = \frac{\left(\frac{NA^2}{\sigma^2}\right)^2}{2N} = N \frac{A^4}{2\sigma^4}$$

$$= \frac{N}{2} \eta_{ED}^2$$

FOR SAME PERFORMANCE  $d_{MF}^2 = d_{ED}^2$   
 $= d^2$

$$\Rightarrow \eta_{MF} = d^2/N \propto 1/N \quad (3 \text{ dB FOR } N \text{ DOUBLING})$$

$$\eta_{ED} = \sqrt{\frac{2d^2}{N}} \propto 1/\sqrt{N} \quad (1.5 \text{ dB FOR } N \text{ DOUBLING})$$



MATCHED FILTER COHERENTLY

COMBINES DATA WHILE ENERGY

DETECTOR INCOHERENTLY COMBINES  
DATA

GLRT FOR COMMON PROBLEMS

1) UNKNOWN AMPLITUDE

$$H_0: x(n) = w(n)$$

$$H_1: x(n) = A s(n) + w(n)$$

↑ ↑ KNOWN

UNKNOWN  $-\infty < A < \infty$

NO UMP TEST SINCE A CAN BE  
POSITIVE OR NEGATIVE

$$L_G(x) = \frac{p(x; \hat{A}, H_1)}{p(x; H_0)}$$

$$p(x; A, H_1) = \frac{1}{(2\pi\sigma^2)^{NT/2}} e^{-\frac{1}{2\sigma^2} \sum_n (x(n) - A s(n))^2}$$

TO FIND MLE OF  $\hat{A}$  MINIMIZE

$$J(A) = \sum_n (x(n) - A s(n))^2$$

$$\Rightarrow \hat{A} = \frac{\sum_{n=0}^{N-1} x(n) s(n)}{\sum_{n=0}^{N-1} s^2(n)}$$

$$\hat{A} = ? \quad \text{IF } S(n) = 1 ?$$

$$\text{LN LG}(x) = -\frac{1}{2\sigma^2} \left( \sum (x(n) - \hat{A}S(n))^2 - \sum x^2(n) \right)$$

$$= -\frac{1}{2\sigma^2} \left( -2\hat{A} \sum x(n)S(n) + \hat{A}^2 \sum S^2(n) \right)$$

$$= \frac{\hat{A}}{\sigma^2} \underbrace{\sum x(n)S(n)}_{\hat{A} \sum S^2(n)} - \frac{1}{2\sigma^2} \hat{A}^2 \sum S^2(n)$$

$$= \hat{A}^2 \frac{\sum S^2(n)}{2\sigma^2} > \text{LN } \gamma$$

$$\text{OR } \hat{A}^2 > \frac{2\sigma^2 \text{LN } \gamma}{\sum_n S^2(n)}$$

$$\text{OR } |\hat{A}| > \sqrt{\frac{2\sigma^2 \text{LN } \gamma}{\sum_n S^2(n)}}$$

$$\text{OR } \left( \sum_{n=0}^{N-1} x(n)S(n) \right)^2 > \gamma'$$

REPLICA-CORRELATOR + SQUARED TO  
ACCOUNT FOR LACK OF KNOWLEDGE  
OF SIGN OF A.

PERFORMANCE GIVEN IN TEXT -  
 ONLY SLIGHT DEGRADATION  
 (SEE FIG 7.3)

## 2) UNKNOWN ARRIVAL TIME

GIVEN A SIGNAL  $s(n)$   $n=0, 1, \dots, M-1$   
 THAT ARRIVES AT  $n=n_0$  OR  $s(n-n_0)$   
 WE WISH TO DETECT IT WITHOUT  
 KNOWING  $n_0$ .

$$H_0: x(n) = w(n) \quad n=0, 1, \dots, N-1$$

$$H_1: x(n) = s(n-n_0) + w(n) \quad "$$

$s(n)$  IS KNOWN AND  $w(n)$  IS WGN,  
 ASSUME ENTIRE SIGNAL IS INCLUDED  
 IN OBSERVATION INTERVAL  $[0, N-1]$

A GLRT DECIDES  $H_1$  IF

$$\frac{p(x; \hat{n}_0, H_1)}{p(x; H_0)} > \gamma$$

NEED MLE OF  $n_0$



$$\begin{aligned}
 p(x; n_0) &= \prod_{n=0}^{n_0-1} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2} x^2/n} \\
 &\cdot \prod_{n=n_0}^{n_0+M-1} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2} (\alpha/n - \beta/n - n_0)^2} \\
 &\cdot \prod_{n=n_0+M}^{N-1} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2} x^2/n}
 \end{aligned}$$

TO MAXIMIZE OVER  $n_0$  NEED TO MAXIMIZE

$$\begin{aligned}
 p(x; n_0) &= \frac{1}{(2\pi\sigma^2)^{N/2}} e^{-\frac{1}{2\sigma^2} \sum_{n=0}^{N-1} x^2/n} \\
 &\cdot \prod_{n=n_0}^{n_0+M-1} e^{-\frac{1}{2\sigma^2} (-2x/n)\beta/n - n_0 + \beta^2/n - n_0^2)}
 \end{aligned}$$

OR MAXIMIZE

$$e^{-\frac{1}{2\sigma^2} \sum_{n=n_0}^{n_0+M-1} (-2x/n)\beta/n - n_0 + \beta^2/n - n_0^2)}$$

OR MINIMIZE

$$\sum_{n=n_0}^{n_0+M-1} (-2x/n)\beta/n - n_0 + \beta^2/n - n_0^2$$

$$\text{BUT } \sum_{n=n_0}^{n_0+M-1} \beta^2/n - n_0^2 = \sum_{n=0}^{M-1} \beta^2/n = \epsilon$$

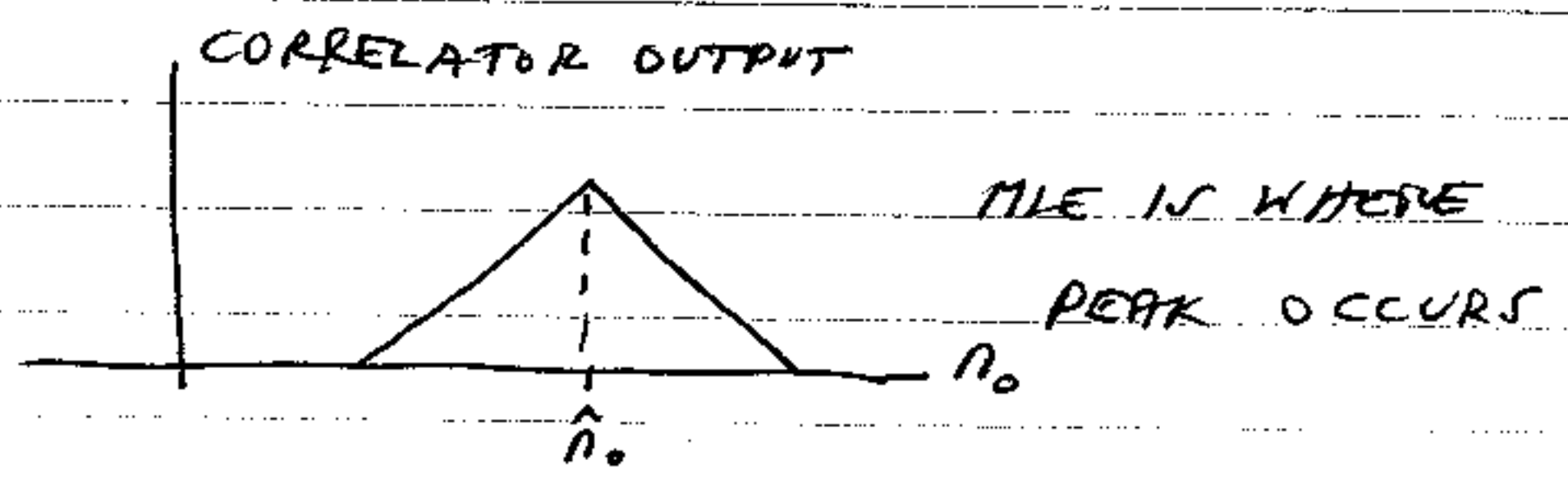
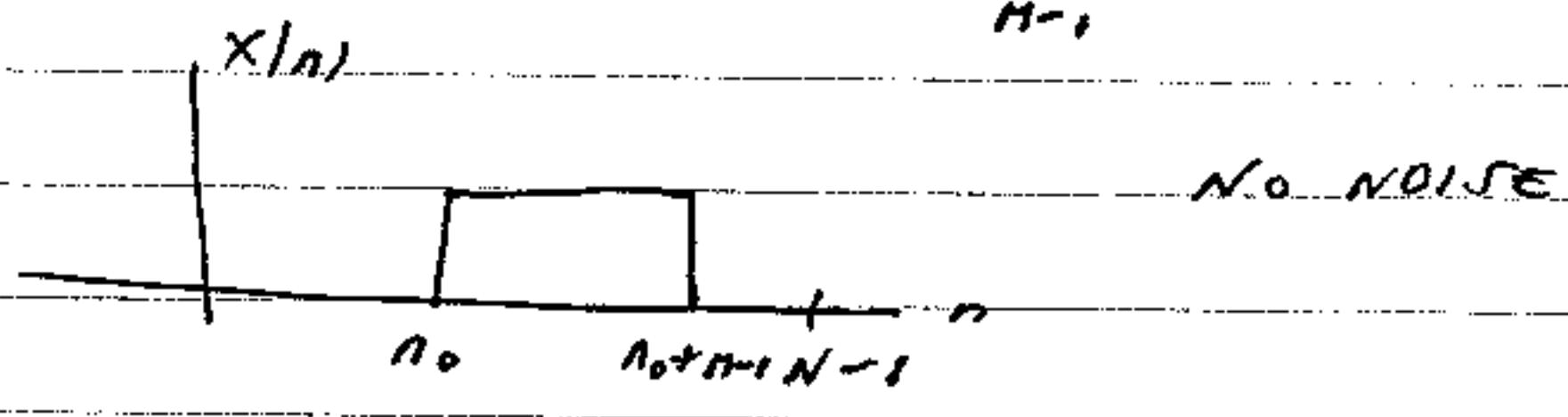
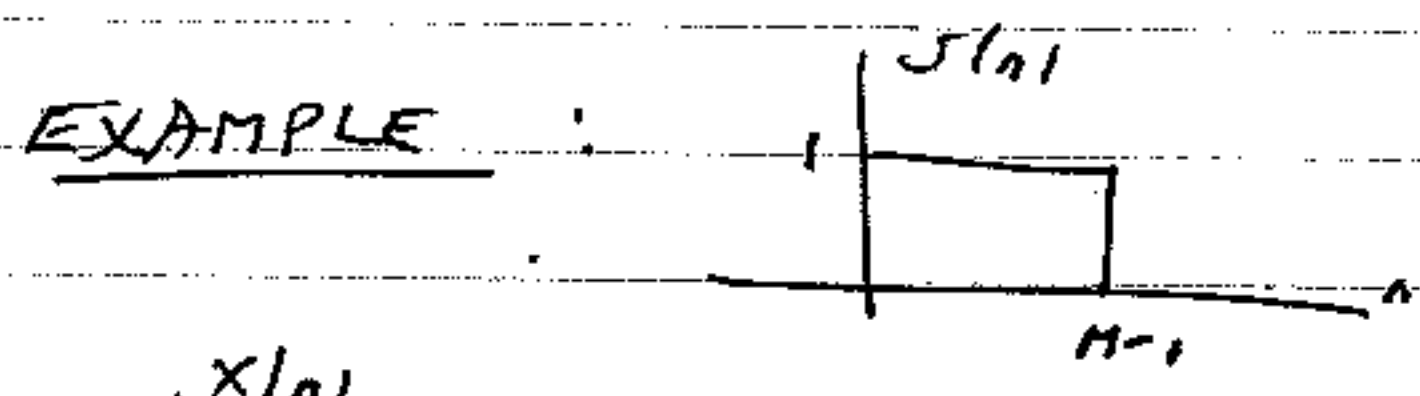
DOES NOT DEPEND ON  $n_0$ .

⇒ MAXIMIZE OVER  $n_0$

$$\sum_{n=n_0}^{n_0+M-1} x(n) s(n-n_0)$$

THIS IS A CORRELATOR THAT CORRELATES AGAINST ALL POSSIBLE SIGNALS - DIFFERENT ARRIVAL TIMES.

⇒ MLE = "RUNNING" CORRELATOR



$$\frac{p(x; \hat{n}_0, H_1)}{p(x; H_0)} = \prod_{n=\hat{n}_0}^{\hat{n}_0+M-1} e^{-\frac{1}{2\sigma^2} (-2x(n)s(n-\hat{n}_0) + s^2(n-\hat{n}_0))}$$

TAKE LOG  $\Rightarrow$  WE DECIDE  $H_0$  IF

$$-\frac{1}{2\sigma^2} \sum_{n=\hat{n}_0}^{\hat{n}_0+M-1} (-2x(n)\sigma/n - \hat{n}_0) + \sigma^2(n - \hat{n}_0) > \text{LNT}$$

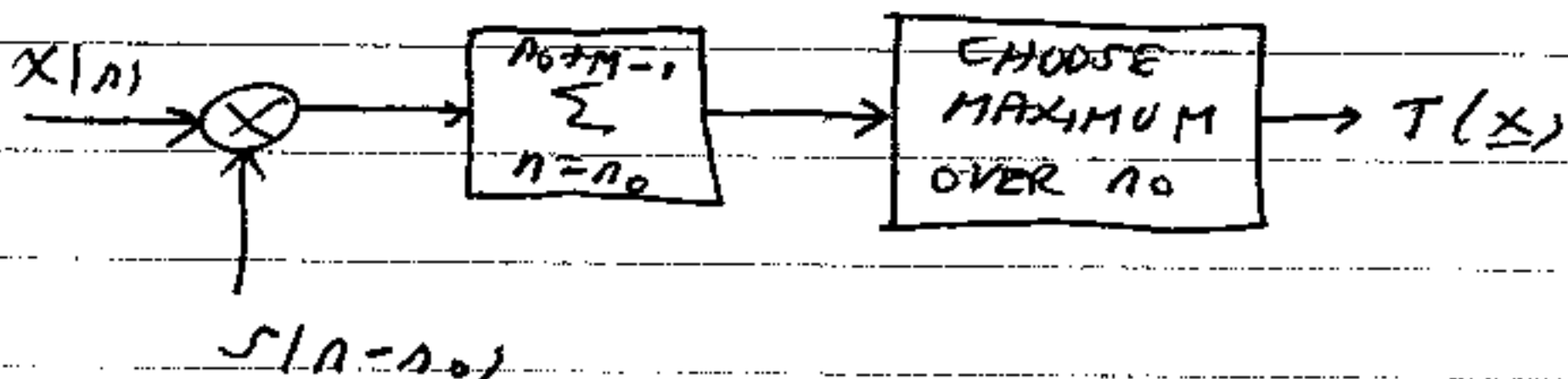
BUT  $\sum_{n=\hat{n}_0}^{\hat{n}_0+M-1} \sigma^2(n - \hat{n}_0) = \Sigma$

$$\sum_{n=\hat{n}_0}^{\hat{n}_0+M-1} x(n)\sigma/n > \sigma^2 \text{LNT} + \Sigma/2 = \gamma'$$

OR FINALLY SINCE  $\hat{n}_0$  IS VALUE THAT MAXIMIZES  $\sum_{n=n_0}^{\hat{n}_0+M-1} x(n)\sigma/n$

WE DECIDE  $H_1$  IF

$$\text{MAX}_{n_0} \sum_{n=n_0}^{\hat{n}_0+M-1} x(n)\sigma/n > \gamma'$$



NOTE THAT IF WE EXCEED THRESHOLD AND DECLARE  $H_1$  (SIGNAL PRESENT), THEN ALSO HAVE MLE OF ARRIVAL TIME OR  $\hat{n}_0$ .

PFA, PD HARD TO FIND ANALYTICALLY  
 NEED PDF OF MAXIMUM OF CORRELATED  
 (DUE TO OVERLAP FOR  $n_0, n_0+1, \dots$ )  
 GAUSSIAN RANDOM VARIABLES.

### SINUSOIDAL DETECTION

IMPORTANT PRACTICAL PROBLEM. WILL  
 CONSIDER

$$H_0: x(n) = w(n) \quad n = 0, 1, \dots, N-1$$

$$H_1: x(n) = A \cos(2\pi f_0 n + \phi) + w(n), "$$

$w(n)$  IS WGN AND SIGNAL IS  
 DETERMINISTIC WITH UNKNOWN  
 PARAMETERS

CASES:

- 1) A UNKNOWN
- 2) A,  $\phi$  UNKNOWN
- 3) A,  $\phi$ ,  $f_0$  UNKNOWN
- \* 4) A,  $\phi$ ,  $f_0$  AND ARRIVAL TIME  
 UNKNOWN

\* DEFER FOR NOW

1) A UNKNOWN

FROM PREVIOUS WORK (SLIDES

153, 154)

DECIDE  $H_1$  IF

$$\left( \sum_{n=0}^{N-1} x[n] s[n] \right)^2 > \gamma'$$

$$\left( \sum_n x[n] \cos(2\pi f_0 n + \phi) \right)^2 > \gamma'$$

NOTE THAT MLE OF A IS

$$\hat{A} = \frac{\sum x[n] s[n]}{\sum s^2[n]}$$

$$= \frac{\sum x[n] \cos(2\pi f_0 n + \phi)}{\sum \cos^2(2\pi f_0 n + \phi)}$$

$$\approx \frac{2}{N} \sum_{n=0}^{N-1} x[n] \cos(2\pi f_0 n + \phi)$$

2) A,  $\phi$  UNKNOWN

$$s[n] = A \cos(2\pi f_0 n + \phi)$$

ASSUME  $A > 0$ , ELSE CAN'T IDENTIFY

A,  $\phi$  UNIQUELY SINCE